

# Categorical Data Analysis and Visualisation

## Part III: Multi-way Contingency Tables

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## Multiple Correspondence Analysis

- Indicator table
- Burt table

## Multiple Correspondence Analysis

So far we have confined our attention to the case where we are graphically summarising the association between two categorical variables.

In the case where we have three categorical variables, we focus on understanding the association between the rows, columns and tubes.

For more than three categorical variables, the depiction can become very messy for categorical visualisation.

When we have multiple categorical variables, we can perform a correspondence analysis by coding the variables in a number of ways. We can

- derive the *indicator table* form of the data
- derive the *Burt table* form of the data
- perform a *joint correspondence analysis, stacking and concatenation*

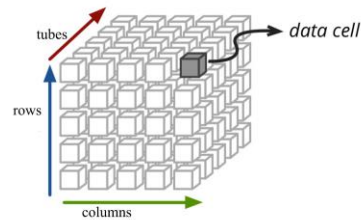
By recoding the contingency table, we perform *multiple correspondence analysis* (MCA)

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## Notation

For a sample size,  $n$ , we denote a three-way contingency table,  $\mathbf{N}$ , so that

- There are  $I$  row categories,  $J$  column categories and  $K$  tube categories
- $\mathbf{N}$  is of size  $I \times J \times K$
- $n_{ijk}$  is the  $(i, j, k)$ 'th cell frequency
- $n_{i..}$  is the  $i$ 'th row marginal frequency
- $n_{.j.}$  is the  $j$ 'th column marginal frequency
- $n_{...k}$  as the  $k$ 'th tube marginal frequency
- the  $(i, j, k)$ 'th cell proportion by  $p_{ijk} = n_{ijk}/n$
- the  $i$ 'th row marginal proportion by  $p_{i..} = \sum_{j=1}^J \sum_{k=1}^K p_{ijk}$
- the  $j$ 'th column marginal proportion by  $p_{.j.} = \sum_{i=1}^I \sum_{k=1}^K p_{ijk}$
- the  $k$ 'th tube marginal proportion by  $p_{...k} = \sum_{i=1}^I \sum_{j=1}^J p_{ijk}$



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## Indicator Matrix

Any sized contingency table can be coded using an *indicator table*, denoted by  $\mathbf{Z}$ .

Suppose we wish to perform a correspondence analysis on our two-way table using its indicator matrix,  $\mathbf{Z}$ .

$\mathbf{Z}$  is formed by concatenating two sub-matrices (one for each variable) such that

$$\mathbf{Z} = [\mathbf{Z}_1 \ \mathbf{Z}_2]$$

Each row of the indicator matrix represents how each individual is classified into the categories.

For the first variable,  $\mathbf{Z}_1$  consists of only the elements 1 and 0; 1 where the individuals has a characteristic, 0 where it doesn't.

$$\mathbf{Z}_1 = \begin{pmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \end{pmatrix} \quad \mathbf{Z}_2 = \begin{pmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \end{pmatrix}$$

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## Indicator Matrix

Therefore, for an  $\mathbf{M}$ -way contingency table of size  $I_1 \times I_2 \times \dots \times I_M \dots$

... that cross classifies  $n$  individuals or units, its indicator table form is of dimension

$$n \times \left( \sum_{m=1}^M I_m \right)$$

For example, an  $I \times J$  contingency table has an indicator table of dimension

$$n \times (I + J)$$

One way to perform multiple correspondence analysis is to perform simple correspondence analysis on the indicator table.

$$\mathbf{Z}_1 = \begin{pmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \end{pmatrix} \quad \mathbf{Z}_2 = \begin{pmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \end{pmatrix}$$

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## Example: Naples Hospital Data

Multiple Correspondence Analysis

- In 2003 a group of Italian researchers investigated Patient Satisfaction (I) in a Naples hospital (n = 1049). The two most important factors
  - Cleanliness of the hospital (J)
  - Management of the hospital (K)

*“How are these three variables related to one another?”*

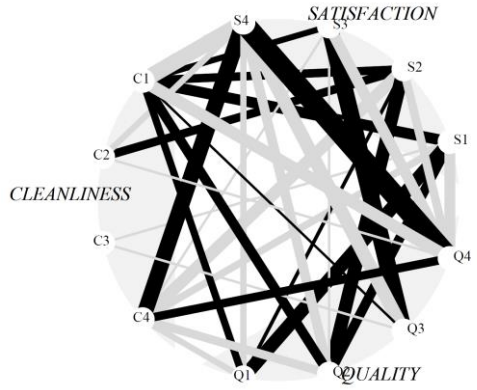
Satisfaction	Cleanliness			
	1 (Poor)	2	3	4 (Exce)
<i>Quality of Management 1 (Low)</i>				
1 (Poor)	16	2	2	1
2 (Fair)	8	1	2	2
3 (Good)	1	4	2	1
4 (Excellent)	1	1	2	3
<i>Quality of Management 2</i>				
1	14	6	2	1
2	26	2	10	6
3	4	1	6	4
4	2	2	2	2
<i>Quality of Management 3</i>				
1	6	2	1	1
2	4	14	2	0
3	36	4	14	82
4	4	8	2	14
<i>Quality of Management 4 (High)</i>				
1	4	2	0	4
2	16	22	10	4
3	44	22	24	50
4	10	24	98	382

7 Source: D' Ambra, L. et al. (2004) In : [naples.dat](http://naples.dat)

## Example: Naples Hospital Data

Multiple Correspondence Analysis

*Cobweb Diagram*



- By using standardised residuals
- Black line indicates *positive* association
  - Grey line indicates a *negative* association.
  - Thick lines indicate a strong association\*.
  - Thin lines weak association\*.

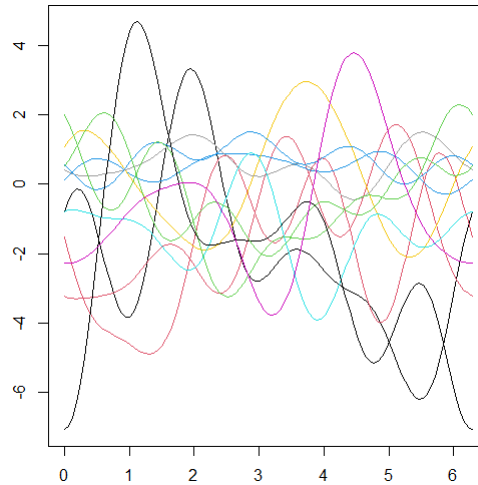
\*Positive and negative association

8 Using the R function `cobweb()` that appears in the Appendix of Upton (2017)



## Example: Naples Hospital Data

Andrew's Plot



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Using the R function `andrewsplot()` in the `pracma` package



## Example: Naples Hospital Data

The Indicator table

```
> indicator.exe(naples.dat)[c(1,10,20,30,40,50,100,150,300,500),]
```

	S1	S2	S3	S4	C1	C2	C3	C4	Q1	Q2	Q3	Q4
Individual 1	1	0	0	0	1	0	0	0	1	0	0	0
Individual 10	1	0	0	0	1	0	0	0	1	0	0	0
Individual 20	1	0	0	0	1	0	0	0	0	1	0	0
Individual 30	1	0	0	0	1	0	0	0	0	1	0	0
Individual 40	1	0	0	0	1	0	0	0	0	0	0	1
Individual 50	1	0	0	0	0	1	0	0	0	0	1	0
Individual 100	0	1	0	0	1	0	0	0	0	0	1	0
Individual 150	0	1	0	0	0	1	0	0	0	0	0	1
Individual 300	0	0	1	0	0	1	0	0	0	0	0	1
Individual 500	0	0	0	1	1	0	0	0	0	0	0	1

$Z_1$

$Z_2$

$Z_3$

Satisfaction

Cleanliness

Quality

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Function `indicator.exe()` appears at the end of these slides (Appendix A)



**Example: Naples Hospital Data**

Multiple Correspondence Analysis

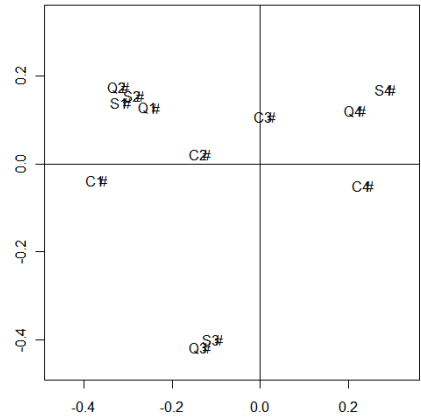
*The Indicator table*

```
> Cav.naples.indicator <-
+   CAvariants(indicator.exe(naples.dat))
> Cav.naples.indicator
.
.
Total inertia 3
```

The inertia values, their percentage contribution to the total inertia and the cumulative percent inertias

	inertia	inertiapc	cuminertiapc
value1	0.6800	22.667	22.667
value2	0.4830	16.100	38.767
value3	0.4045	13.483	52.250
value4	0.3478	11.593	63.843
value5	0.3242	10.807	74.650
value6	0.2643	8.811	83.461
value7	0.2082	6.941	90.402
value8	0.1746	5.821	96.223
value9	0.1133	3.777	100.000

$$\text{Total Inertia} = \frac{I+J+K}{3} - 1 = 3$$



... poor quality (38.8%) 2D visual summary of the association ...

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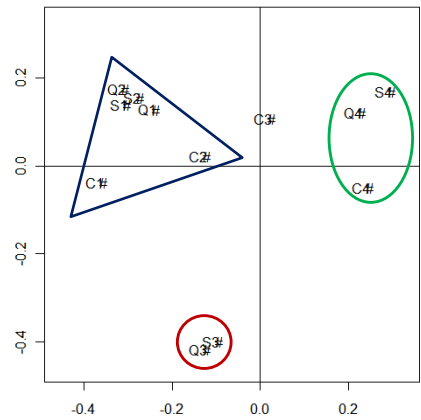
**Example: Naples Hospital Data**

Multiple Correspondence Analysis

*The Indicator table*

There are three main clusters of points:

- 1) There exists an association between the two lowest levels of *quality*, *satisfaction* and cleanliness
- 2) Good quality of management is strongly associated with good satisfaction
- 3) There is an association between excellent levels of satisfaction, quality of management and cleanliness



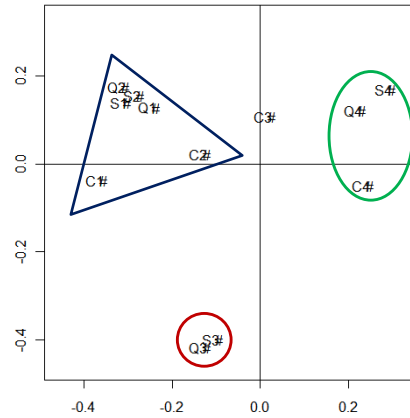
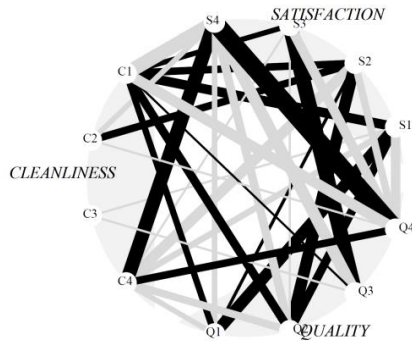
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## Example: Naples Hospital Data

Multiple Correspondence Analysis

*The Indicator table*



- Correspondence plot may only reflect 38.8% of the association but it is consistent with the cobweb diagram
- Correspondence plot shows how categories of the same variable compare
- Cobweb diagram looks rather messy



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## Burt Matrix

Multiple Correspondence Analysis

For the coding of a two-way contingency table, the *Burt table* consists of the concatenation of the original contingency table with diagonal matrices consisting of the row, column and tube marginal frequencies

The Burt table (Burt, 1950) for a three-way table is

$$\begin{aligned}
 \mathbf{B} &= \begin{pmatrix} \mathbf{D}_1 & \mathbf{N}_{12} & \mathbf{N}_{13} \\ \mathbf{N}_{12}^T & \mathbf{D}_J & \mathbf{N}_{22} \\ \mathbf{N}_{13}^T & \mathbf{N}_{23}^T & \mathbf{D}_K \end{pmatrix} \\
 &= \begin{pmatrix} \mathbf{Z}_1^T \mathbf{Z}_1 & \mathbf{Z}_1^T \mathbf{Z}_2 & \mathbf{Z}_1^T \mathbf{Z}_3 \\ \mathbf{Z}_2^T \mathbf{Z}_1 & \mathbf{Z}_2^T \mathbf{Z}_2 & \mathbf{Z}_2^T \mathbf{Z}_3 \\ \mathbf{Z}_3^T \mathbf{Z}_1 & \mathbf{Z}_3^T \mathbf{Z}_2 & \mathbf{Z}_3^T \mathbf{Z}_3 \end{pmatrix}
 \end{aligned}$$



Sir Cyril Burt  
(1883 – 1971)

where, for example,  $\mathbf{N}_{12}$  is the two-way contingency table of the first (row) and second (column) variables formed by aggregating across the third (tube) variable

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## Example: Naples Hospital Data

### The Burt table

Multiple Correspondence Analysis

```
> burt.exe(naples.dat)
  S1  S2  S3  S4  C1  C2  C3  C4  Q1  Q2  Q3  Q4
S1  64  0  0  0  40  12  5  7  21  23  10  10
S2  0  12  DI 0  0  54  39  24  12  13  44  20  52
S3  0  0  0  9  0  85  31  46  137  8  15  136  140
S4  0  0  0  0  557  17  35  104  401  7  8  28  514
C1  40  54  85  17  196  0  0  0  26  46  50  74
C2  12  39  31  35  0  11  DJ 0  0  8  11  28  70
C3  5  24  46  104  0  0  0  9  0  8  20  19  132
C4  7  12  137  401  0  0  0  557  7  13  97  440
Q1  21  13  8  7  26  8  8  7  49  0  0  0
Q2  23  44  15  8  46  11  20  13  0  9  DK 0
Q3  10  20  136  28  50  28  19  97  0  0  0  0
Q4  10  52  140  514  74  70  132  440  0  0  0  716
```

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Function `burt.exe()` appears at the end of these slides (Appendix B)

## Example: Naples Hospital Data

### The Burt table

Multiple Correspondence Analysis

```
> burt.exe(naples.dat)
  S1  S2  S3  S4  C1  C2  C3  C4  Q1  Q2  Q3  Q4
S1  64  0  0  0  40  12  5  7  21  23  10  10
S2  0  12  9  0  0  54  39  24  12  13  44  20  52
S3  0  0  0  299  0  85  31  46  137  8  15  136  140
S4  0  0  0  0  557  17  35  104  401  7  8  28  514
C1  40  54  85  17  196  0  0  0  26  46  50  74
C2  12  39  31  35  0  117  0  0  8  11  28  70
C3  5  24  46  104  0  0  179  0  8  20  19  132
C4  7  12  137  401  0  0  0  557  7  13  97  440
Q1  21  13  8  7  26  8  8  7  49  0  0  0
Q2  23  44  15  8  46  11  20  13  0  90  0  0
Q3  10  20  136  28  50  28  19  97  0  0  194  0
Q4  10  52  140  514  74  70  132  440  0  0  0  716
```

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Function `burt.exe()` appears at the end of these slides (Appendix B)



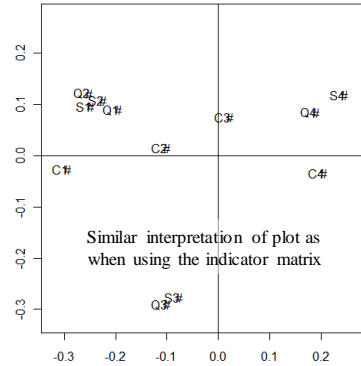
## Example: Naples Hospital Data

### The Burt table

```
> CAvariant.naples.burt <- CAvariants(burt.exe(naples.dat))
> CAvariant.naples.burt
.
.
.
Total inertia 1.242
```

The inertia values, their percentage contribution to the total inertia and the cumulative percent inertias

	inertia	inertiapc	cuminertiapc
value1	0.462	37.234	37.234
value2	0.233	18.787	56.020
value3	0.164	13.171	69.191
value4	0.121	9.742	78.933
value5	0.105	8.464	87.397
value6	0.070	5.625	93.023
value7	0.043	3.489	96.512
value8	0.030	2.454	98.966
value9	0.013	1.034	100.000
value10	0.000	0.000	100.000
value11	0.000	0.000	100.000



- OK quality correspondence plot (56.0%)
- Better than when using the indicator matrix (38.8%)



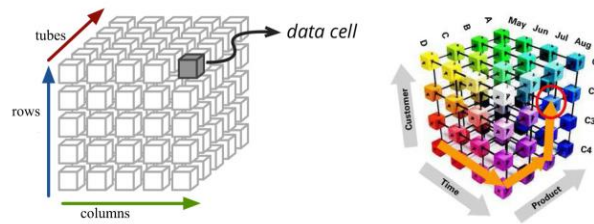
## Multi-way Correspondence Analysis & Symmetric Association

- Partitioning Pearson's Statistic,
- Tucker3 Decomposition,
- Interactive Biplots

## Multi-way Correspondence Analysis

An alternative approach to performing correspondence analysis is to consider the following

- Rather than recoding a three- (or multi-) way contingency table into a two way matrix form then performing simple correspondence analysis on that recoded matrix, instead
- Treat the three-way (say) contingency table as a cube (like a Rubic's cube)



This variant is called *multi-way correspondence analysis* (MWCA) . . . and . . .  
 . . . requires a three- (multi-) way analogue of singular value decomposition.

## Pearson's Chi-squared Statistic

For three *symmetrically* associated categorical variables that are cross-classified to form a three-way contingency table define the *Pearson residuals* by

$$\pi_{ijk} = \frac{p_{ijk}}{p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}} - 1$$

Then Pearson's three-way chi-squared statistic is

$$\begin{aligned} X^2 &= \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K \frac{(p_{ijk} - p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k})^2}{p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}} \\ &= \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k} \left( \frac{p_{ijk}}{p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}} - 1 \right)^2 \\ &= \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k} \pi_{ijk}^2 \end{aligned}$$

## Decomposing Pearson's Residuals

- To formally test whether there is an (symmetric) association between the row, column, and tubes variables we perform a *chi-squared test of independence*
- Pearson's chi-squared statistic is . . .

$$X^2 = n \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K \frac{(p_{ijk} - p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k})^2}{p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}}$$

. . . and has a chi-squared random variable with

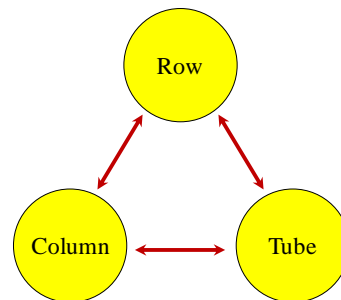
$$(I - 1)(J - 1) + (I - 1)(K - 1) + (J - 1)(K - 1) + (I - 1)(J - 1)(K - 1)$$

degrees of freedom.

## Three-way Symmetric Association

### Pearson's three-way chi-squared statistic

- Row (I), Column (J) and Tube (K) Variables



$$\pi_{ijk} = \frac{p_{ijk}}{p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}} - 1$$

$$X^2 = n \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k} \pi_{ijk}^2$$

## Decomposing Pearson's Residuals

For three *symmetrically* associated categorical variables that are cross-classified to form a three-way contingency table define the *Pearson residuals* by

$$\pi_{ijk} = \frac{p_{ijk}}{p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}} - 1$$

To perform a multi-way correspondence analysis on a three-way contingency table we resort to using two three-way analogues of SVD

*CANDECOMP/PARAFAC Decomposition*

$$\pi_{ijk} = \sum_{m=1}^M a_{im} b_{jm} c_{km} \lambda_m + e_{ijk}$$

- *PARAFAC* (PARAllel FACtor analysis)
    - Developed by Harshman (1970)
  - *CANDECOMP* (CANonical DECOMPosition)
    - Developed by Carroll & Chang (1970)
- } Not easy to define the best rank M!!  
Mathematically equivalent (elaborated upon by many)

## Decomposing Pearson's Residuals

For three *symmetrically* associated categorical variables that are cross-classified to form a three-way contingency table define the *Pearson residuals* by

$$\pi_{ijk} = \frac{p_{ijk}}{p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}} - 1$$

To perform a multi-way correspondence analysis on a three-way contingency table we resort to using two three-way analogues of SVD

*Tucker3 Decomposition*

$$\pi_{ijk} = \sum_{p=1}^P \sum_{q=1}^Q \sum_{r=1}^R \underbrace{a_{ip} b_{jq} c_{kr}}_{\text{Component matrices}} \lambda_{pqr} + \underbrace{e_{ijk}}_{\text{Error}}$$

- Developed by Tucker (1963) . . . elaborated upon by Tucker (1964, 1966)
- See also Kroonenberg (1983, 2008), Paatero & Andersson (1999), Bro & Kiers (2003), Kiers (2004), Pravdova, Estienne, Walczak & Massart (2001), Beh & Lombardo (2014, 2021) and many others

## Decomposing Pearson's Residuals

For three *symmetrically* associated categorical variables that are cross-classified to form a three-way contingency table define the *Pearson residuals* by

$$\pi_{ijk} = \frac{p_{ijk}}{p_{i..}p_{.j.}p_{..k}} - 1$$

To perform a multi-way correspondence analysis on a three-way contingency table we resort to using two three-way analogues of SVD

*Tucker3 Decomposition*

$$\hat{\pi}_{ijk} = \sum_{p=1}^P \sum_{q=1}^Q \sum_{r=1}^R a_{ip} b_{jq} c_{kr} \lambda_{pqr}$$

Choose a P, Q and R that minimises...not easy the solution are not nested (ALS algorithm)...

$$SSE = \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{i..}p_{.j.}p_{..k} (\pi_{ijk} - \hat{\pi}_{ijk})^2$$

## Decomposing Pearson's Residuals

For three *symmetrically* associated categorical variables that are cross-classified to form a three-way contingency table define the *Pearson residuals* by

$$\pi_{ijk} = \frac{p_{ijk}}{p_{i..}p_{.j.}p_{..k}} - 1$$

To perform a multi-way correspondence analysis on a three-way contingency table we resort to using two three-way analogues of SVD

*Tucker3 Decomposition*

$$\hat{\pi}_{ijk} = \sum_{p=1}^P \sum_{q=1}^Q \sum_{r=1}^R a_{ip} b_{jq} c_{kr} \lambda_{pqr}$$

Choose a P, Q and R...

**Deviance Plot:** A measure of *goodness of fit* on y-axis and the degree of freedom on x-axis- *select a model on or close to an elbow in the higher boundary of the convex hull*

(Timmerman, Kiers 2000; Ceulemans, Kiers, 2006; Lombardo, van de Velden, Beh 2022)

## Decomposing Pearson's Residuals

For three *symmetrically* associated categorical variables that are cross-classified to form a three-way contingency table define the *Pearson residuals* by

$$\pi_{ijk} = \frac{p_{ijk}}{p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}} - 1$$

To perform a multi-way correspondence analysis on a three-way contingency table we resort to using two three-way analogues of SVD

*Tucker3 Decomposition*

$$\hat{\pi}_{ijk} = \sum_{p=1}^P \sum_{q=1}^Q \sum_{r=1}^R a_{ip} b_{jq} c_{kr} \lambda_{pqr}$$

*Column-Tube Interactive Biplot* (but you can also define principal coords for rows)

$$\begin{aligned} \tilde{f}_{ip} &= a_{ip} & \tilde{g}_{jkp} &= \sum_{q=1}^Q \sum_{r=1}^R b_{jq} c_{kr} \lambda_{pqr} \\ \text{(rows points)} & & \text{("interactive" column-tube points)} & \end{aligned}$$

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## Decomposing Pearson's Residuals

For three *symmetrically* associated categorical variables that are cross-classified to form a three-way contingency table define the *Pearson residuals* by

$$\pi_{ijk} = \frac{p_{ijk}}{p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}} - 1$$

To perform a multi-way correspondence analysis on a three-way contingency table we resort to using two three-way analogues of SVD

*Tucker3 Decomposition*

$$\hat{\pi}_{ijk} = \sum_{p=1}^P \sum_{q=1}^Q \sum_{r=1}^R a_{ip} b_{jq} c_{kr} \lambda_{pqr}$$

*Row-Tube Interactive Biplot* (but you can also define principal coords for columns)

$$\begin{aligned} \tilde{f}_{jq} &= b_{jq} & \tilde{g}_{ikq} &= \sum_{p=1}^P \sum_{r=1}^R a_{ip} c_{kr} \lambda_{pqr} \\ \text{(columns points)} & & \text{("interactive" row-tube points)} & \end{aligned}$$

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## Decomposing Pearson's Residuals

For three *symmetrically* associated categorical variables that are cross-classified to form a three-way contingency table define the *Pearson residuals* by

$$\pi_{ijk} = \frac{p_{ijk}}{p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}} - 1$$

To perform a multi-way correspondence analysis on a three-way contingency table we resort to using two three-way analogues of SVD

*Tucker3 Decomposition*

$$\hat{\pi}_{ijk} = \sum_{p=1}^P \sum_{q=1}^Q \sum_{r=1}^R a_{ip} b_{jq} c_{kr} \lambda_{pqr}$$

*Row-Column Interactive Biplot* (but you can also define principal coords for tubes)

$$\tilde{f}_{kr} = c_{kr} \quad \tilde{g}_{ijq} = \sum_{p=1}^P \sum_{q=1}^Q a_{ip} b_{jq} \lambda_{pqr}$$

(tubes points)                      (“interactive” row-column points)

## Example: Sustainable Development Goals Data

The cross-classification of the three categorical variables (RES, BCA & GEO) that produce the  $4 \times 4 \times 6$  contingency table . . .

- RES: indicator of renewable energy share of total final energy consumption
- BCA: an indicator of adjusted emission growth rate for black carbon
- GEO: geographical area

	RES	BCA			
		(0, 8.13]	(8.13, 23.1]	(23.1, 49.7]	(49.7, 96.4]
		BCA1	BCA2	BCA3	BCA4
<b>Africa</b>					
(0, 34.5] RES1	1	0	4	8	
(34.5, 53.7] RES2	0	2	4	19	
(53.7, 80.7] RES3	3	2	1	4	
(80.7, 100] RES4	1	1	0	1	
<b>America</b>					
RES1	1	0	6	2	
RES2	0	2	0	1	
RES3	0	3	1	0	
RES4	1	3	2	1	
<b>Asia</b>					
RES1	5	1	2	0	
RES2	6	2	3	1	
RES3	5	3	3	3	
RES4	4	2	0	0	
<b>Australia</b>					
RES1	1	0	0	0	
RES2	0	1	2	0	
RES3	1	1	5	0	
RES4	1	0	0	0	
<b>Caribbean</b>					
RES1	3	2	0	0	
RES2	1	1	0	1	
RES3	0	0	0	0	
RES4	2	3	0	0	
<b>Europe</b>					
RES1	1	0	0	0	
RES2	0	0	1	0	
RES3	0	7	5	0	
RES4	3	11	7	3	

**Example: Sustainable Development Goals Data**

Multi-Way Correspondence Analysis

The cross-classification of the three categorical variables (RES, BCA & GEO) that produce the  $4 \times 4 \times 6$  contingency table . . .

Chi-squared test of independence:

- $X^2 = 290.035$
- A p-value  $< 0.001$
- Thus, a statistically significant association exists between at least two of the variables

. . . but . . .

there are other features we'll talk about soon

RES		BCA			
		(0, 8.13]	(8.13, 23.1]	(23.1, 49.7]	(49.7, 96.4]
		BCA1	BCA2	BCA3	BCA4
<b>Africa</b>	(0, 34.5] RES1	1	0	4	8
	(34.5, 53.7] RES2	0	2	4	19
	(53.7, 80.7] RES3	3	2	1	4
	(80.7, 100] RES4	1	1	0	1
<b>America</b>	RES1	1	0	6	2
	RES2	0	2	0	1
	RES3	0	3	1	0
	RES4	1	3	2	1
<b>Asia</b>	RES1	5	1	2	0
	RES2	6	2	3	1
	RES3	5	3	3	3
	RES4	4	2	0	0
<b>Australia</b>	RES1	1	0	0	0
	RES2	0	1	2	0
	RES3	1	1	5	0
	RES4	1	0	0	0
<b>Carribean</b>	RES1	3	2	0	0
	RES2	1	1	0	1
	RES3	0	0	0	0
	RES4	2	3	0	0
<b>Europe</b>	RES1	1	0	0	0
	RES2	0	0	1	0
	RES3	0	7	5	0
	RES4	3	11	7	3

**Example: Sustainable Development Goals Data**

Multi-Way Correspondence Analysis

The cross-classification of the three categorical variables (RES, BCA & GEO) that produce the  $4 \times 4 \times 6$  contingency table . . .

Chi-squared test of independence:

- $X^2 = 290.035$

$$X^2 = X_{IJ}^2 + X_{IK}^2 + X_{JK}^2 + X_{IJK}^2$$

Two-way interactions

Three-way interactions

following partition of Lancaster (1951)

RES		BCA			
		(0, 8.13]	(8.13, 23.1]	(23.1, 49.7]	(49.7, 96.4]
		BCA1	BCA2	BCA3	BCA4
<b>Africa</b>	(0, 34.5] RES1	1	0	4	8
	(34.5, 53.7] RES2	0	2	4	19
	(53.7, 80.7] RES3	3	2	1	4
	(80.7, 100] RES4	1	1	0	1
<b>America</b>	RES1	1	0	6	2
	RES2	0	2	0	1
	RES3	0	3	1	0
	RES4	1	3	2	1
<b>Asia</b>	RES1	5	1	2	0
	RES2	6	2	3	1
	RES3	5	3	3	3
	RES4	4	2	0	0
<b>Australia</b>	RES1	1	0	0	0
	RES2	0	1	2	0
	RES3	1	1	5	0
	RES4	1	0	0	0
<b>Carribean</b>	RES1	3	2	0	0
	RES2	1	1	0	1
	RES3	0	0	0	0
	RES4	2	3	0	0
<b>Europe</b>	RES1	1	0	0	0
	RES2	0	0	1	0
	RES3	0	7	5	0
	RES4	3	11	7	3



## Example: Sustainable Development Goals Data

Multi-Way Correspondence Analysis

Partitioning Pearson's Three-way Chi-squared Statistic

Component	$X_{ij}^2$	$X_{ik}^2$	$X_{jk}^2$	$X_{ijk}^2$	$X^2$
Term	34.970	82.816	95.677	76.573	290.035
P-value	<0.001	<0.001	<0.001	0.002	<0.001
% Contribution	12%	29%	33%	26%	100%

- There is a statistically significant association between at least two of the variables
- There is a statistically significant association between each **pair** of the three variables
- There is a statistically significant association between **all** three variables . . . but . . .
- The association is dominated by the association between BCA & GEO (33%) followed by RES & BCA (29%)

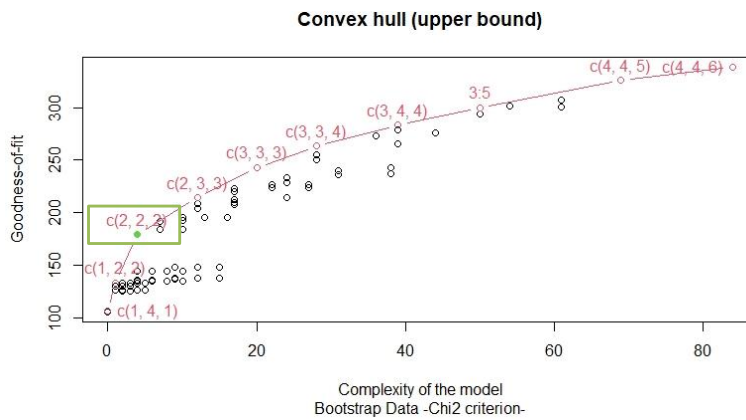


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## Example: Sustainable Development Goals Data

Multi-Way Correspondence Analysis

```
tunelocal(sustainable.dat, boots = T, boottype = "bootpsimple", nboots= 1000)
```



**Deviance Plot** Goodness of Fit = Chi-squared (vertical axis), df (horizontal axis).

Model D222 is on the hull: 
$$st(i) = \frac{\left(\frac{dev(i+1) - dev(i)}{df(i+1) - df(i)}\right)}{\left(\frac{dev(i) - dev(i-1)}{df(i) - df(i-1)}\right)}$$



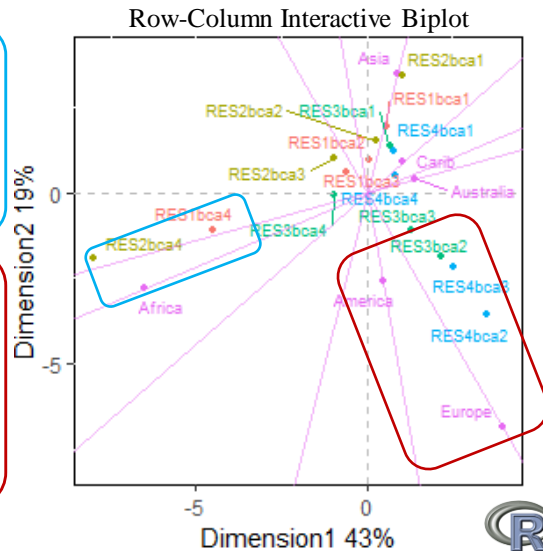
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## Example: Sustainable Development Goals Data

```
> res.ca3 <- CA3variants(sustainable.dat, dims = c(2,2,2), ca3type = "CA3")
> plot(res.ca3, biptype = "row-column", scaleplot = 1.5)
```

- Association between the lowest levels of *RES* and the highest level of black carbon (RES2bca4 – RES1bca4) in Africa

- The highest levels of *RES* are strongly associated with middle-low levels of black-carbon (RES4bca2-RES4bca3) in Europe and America

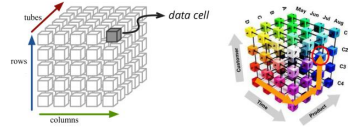


## Multi-way Correspondence Analysis & Asymmetric Association

- Marcotorchino's Index
- Gray-Williams Index
- Lombardo/Lambda Index

## Three-way Analogs of $\tau_{A|B}$

For three asymmetrically associated variables:  
dependent on the structure of the asymmetry



Marcotorchino Index

*Marcotorchino Index*

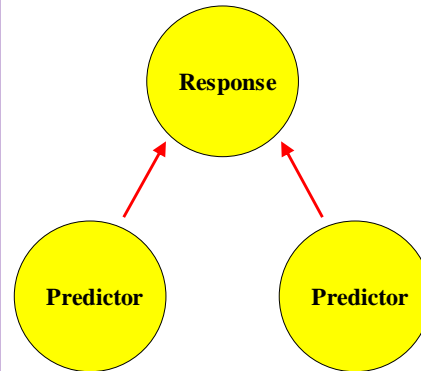
$$\tau_M = \frac{\sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{\bullet j \bullet} p_{\bullet \bullet k} \pi_{ijk}^2}{1 - \sum_{i=1}^I p_{i \bullet \bullet}^2}$$

where

$$\pi_{ijk} = \frac{p_{ijk}}{p_{\bullet j \bullet} p_{\bullet \bullet k}} - p_{i \bullet \bullet}$$

- Two predictor variables (assumed/treated **independent**)
- One response variable

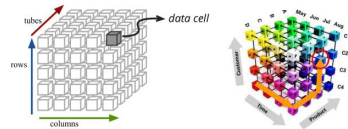
(Marcotorchino, F., 1984)



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## Three-way Analogs of $\tau_{A|B}$

For three asymmetrically associated variables:  
dependent on the structure of the asymmetry



Gray-Williams Index

*Gray – Williams Index*

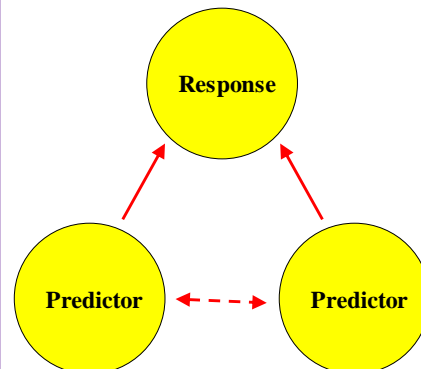
$$\tau_{GW} = \frac{\sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{\bullet j k} \pi_{ijk}^2}{1 - \sum_{i=1}^I p_{i \bullet \bullet}^2}$$

where

$$\pi_{ijk} = \frac{p_{ijk}}{p_{\bullet j k}} - p_{i \bullet \bullet}$$

- Two predictor variables (assumed/treated **symmetrically associated**)
- One response variable

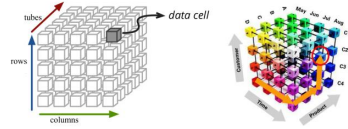
(Gray & Williams, 1981)



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## Three-way Analogs of $\tau_{A|B}$

For three asymmetrically associated variables:  
dependent on the structure of the asymmetry



*Lombardo/Delta Index*

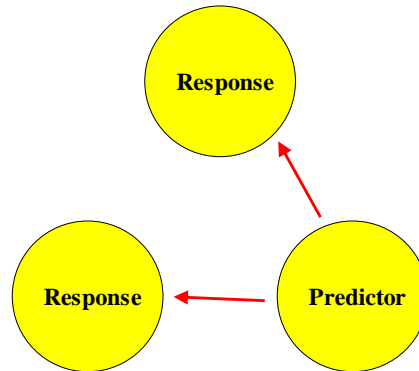
$$\tau_D = \frac{\sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{\bullet jk} \pi_{ijk}^2}{1 - \sum_{i=1}^I p_{i\bullet\bullet}^2}$$

where

$$\pi_{ijk} = \frac{p_{ijk}}{p_{\bullet\bullet k}} - p_{i\bullet\bullet} p_{\bullet j\bullet}$$

- One predictor variable
- Two response variables (assumed/treated **independent**)

(Lombardo, 2011)



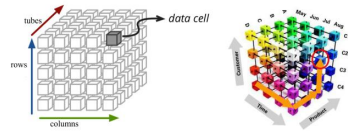
Lombardo/Delta Index

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## Three-way Analogs of $\tau_{A|B}$

For three asymmetrically associated variables:  
dependent on the structure of the asymmetry

Regardless of the  $\pi_{ijk}$  being considered, there  
are two ways in which it can be decomposed  
that are treated as three-way analogues of SVD:



*PARAFAC/CANDECOMP Models*

$$\pi_{ijk} = \sum_{m=1}^M a_{im} b_{jm} c_{km} \lambda_m + e_{ijk}$$

*Tucker3 Decomposition*

$$\pi_{ijk} = \sum_{p=1}^P \sum_{q=1}^Q \sum_{r=1}^R a_{ip} b_{jq} c_{kr} \lambda_{pqr} + e_{ijk}$$

Asymmetric Association

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**Example: Naples Hospital Data**

Multi-way Correspondence Analysis

**Marcotorchino Index**

$$\tau_M = \tau_{IJ} + \tau_{IK} + \tau_{JK} + \tau_{IJK}$$

$$C_{IJ} = 445.614 \quad (\text{Satisfaction vs Cleanliness})$$

$$C_{IK} = 712.186 \quad (\text{Satisfaction vs Management})$$

$$C_{JK} = 212.348 \quad (\text{Cleanliness vs Management})$$

$$C_{IJK} = 299.434 \quad (\text{Three-way Association})$$

$$C_M = 1669.581 \quad (\text{Total Association})$$

		Cleanliness			
		1 (Poor)	2	3	4 (Exce)
response	Quality of Management 1 (Low)				
	1 (Poor)	16	2	2	1
	2 (Fair)	8	1	2	2
	3 (Good)	1	4	2	1
	4 (Excellent)	1	1	2	3
		Quality of Management 2			
	1	14	6	2	1
	2	26	2	10	6
	3	4	1	6	4
	4	2	2	2	2
		Quality of Management 3			
	1	6	2	1	1
	2	4	14	2	0
	3	36	4	14	82
	4	4	8	2	14
		Quality of Management 4 (High)			
	1	4	2	0	4
	2	16	22	10	4
	3	44	22	24	50
	4	10	24	98	382

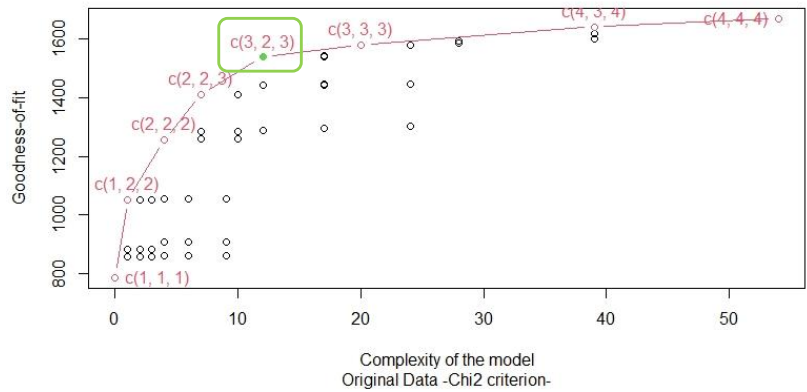
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**Example: Naples Hospital Data**

Multi-way Correspondence Analysis

```
tunelocal(naples.dat, boots = True, bootype = "bootpsimple", nboots= 1000)
```

**Convex hull (upper bound)**



**Deviance Plot** Goodness of Fit = Marcotorchino (vertical axis), df (horizontal axis).

Model D323 is on the hull: 
$$st(i) = \left( \frac{dev(i+1) - dev(i)}{df(i+1) - df(i)} \right) / \left( \frac{dev(i) - dev(i-1)}{df(i) - df(i-1)} \right)$$

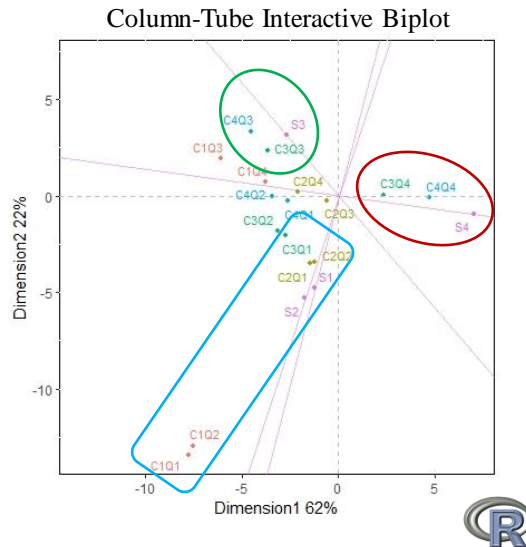


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## Example: Naples Hospital Data

```
> res.ca3 <- CA3variants(naples.dat, dims=c(3,2,3),ca3type = "NSCA3")
> plot(res.ca3, biptype = "pred", scaleplot = 1.5)
```

- Predictability of the two lowest levels of *satisfaction* given quality & cleanliness
- Good satisfaction is well predicted by the good quality of management & cleanliness
- Association between excellent levels of satisfaction given excellent levels of cleanliness & quality of management



## Multi-way Correspondence Analysis & Symmetric Association

- Cressie-Read family of Divergence Statistics

## Three-way Family of Divergence Statistics

For some  $\delta$ , the Cressie-Read family of divergence statistics for a three-way contingency table is

$$CR(\delta) = \frac{2n}{\delta(\delta + 1)} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \left[ \left( \frac{p_{ijk}}{p_{i..} p_{.j.} p_{..k}} \right)^\delta - 1 \right]$$

where  $\delta \in (-\infty, \infty)$ . (Pardo, 1996; Lombardo & Beh, 2022)

This is a chi-squared random variable with

$$(I - 1)(J - 1) + (I - 1)(K - 1) + (J - 1)(K - 1) + (I - 1)(J - 1)(K - 1)$$

degrees of freedom.

## Three-way Family of Divergence Statistics

For some  $\delta$ , the Cressie-Read family of divergence statistics for a three-way contingency table is

$$CR(\delta) = \frac{2n}{\delta(\delta + 1)} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \left[ \left( \frac{p_{ijk}}{p_{i..} p_{.j.} p_{..k}} \right)^\delta - 1 \right]$$

where  $\delta \in (-\infty, \infty)$ . (Pardo, 1996; Lombardo & Beh, 2022)

*Pearson's chi-squared statistic*

$$CR(\delta = 1) = X^2 = n \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K \frac{(p_{ijk} - p_{i..} p_{.j.} p_{..k})^2}{p_{i..} p_{.j.} p_{..k}}$$

*Likelihood Ratio Statistic*

$$CR(0) = G^2 = 2n \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \ln \left( \frac{p_{ijk}}{p_{i..} p_{.j.} p_{..k}} \right)$$

## Three-way Family of Divergence Statistics

For some  $\delta$ , the Cressie-Read family of divergence statistics for a three-way contingency table is

$$CR(\delta) = \frac{2n}{\delta(\delta + 1)} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \left[ \left( \frac{p_{ijk}}{p_{i\cdot\cdot} p_{\cdot j\cdot} p_{\cdot\cdot k}} \right)^\delta - 1 \right]$$

where  $\delta \in (-\infty, \infty)$ . (Pardo, 1996; Lombardo & Beh, 2022)

*Freeman-Tukey statistic*

$$CR\left(-\frac{1}{2}\right) = T^2 = 4n \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K \left( \sqrt{p_{ijk}} - \sqrt{p_{i\cdot\cdot} p_{\cdot j\cdot} p_{\cdot\cdot k}} \right)^2$$

*Cressie-Read Statistic*

$$CR\left(\frac{2}{3}\right) = CR^2 = \frac{9n}{5} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{i\cdot\cdot} p_{\cdot j\cdot} p_{\cdot\cdot k} \left( \left( \frac{p_{ijk}}{p_{i\cdot\cdot} p_{\cdot j\cdot} p_{\cdot\cdot k}} \right)^{2/3} - 1 \right)^2$$

## Partitioning the Family of Divergence Statistics

Lombardo & Beh (2022) showed that the three-way Cressie-Read divergence statistic can be partitioned as follows

$$\begin{aligned} CR(\delta) &= \frac{2n}{\delta(\delta + 1)} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \left[ \left( \frac{p_{ijk}}{p_{i\cdot\cdot} p_{\cdot j\cdot} p_{\cdot\cdot k}} \right)^\delta - 1 \right] \\ &= \frac{2n}{\delta(\delta + 1)} \sum_{i=1}^I \sum_{j=1}^J p_{ij\cdot} \left[ \left( \frac{p_{ijk}}{p_{i\cdot\cdot} p_{\cdot j\cdot}} \right)^\delta - 1 \right] + \frac{2n}{\delta(\delta + 1)} \sum_{i=1}^I \sum_{k=1}^K p_{i\cdot k} \left[ \left( \frac{p_{i\cdot k}}{p_{i\cdot\cdot} p_{\cdot\cdot k}} \right)^\delta - 1 \right] \\ &= \frac{2n}{\delta(\delta + 1)} \sum_{j=1}^J \sum_{k=1}^K p_{\cdot jk} \left[ \left( \frac{p_{\cdot jk}}{p_{\cdot j\cdot} p_{\cdot\cdot k}} \right)^\delta - 1 \right] + \frac{2n}{\delta(\delta + 1)} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \left[ \left( \frac{p_{ijk}}{\alpha p_{i\cdot\cdot} p_{\cdot j\cdot} p_{\cdot\cdot k}} \right)^\delta - 1 \right] \\ &= \boxed{CR_{IJ}(\delta)} + CR_{IK}(\delta) + CR_{JK}(\delta) + CR_{IJK}(\delta) \end{aligned}$$

Cressie-Read family of divergence statistics for only the **row** and **column** variables (aggregating across the **tube** variable)



## Partitioning the Family of Divergence Statistics

Lombardo & Beh (2022) showed that the three-way Cressie-Read divergence statistic can be partitioned as follows

$$\begin{aligned}
 CR(\delta) &= \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \left[ \left( \frac{p_{ijk}}{p_{i..} p_{.j.} p_{..k}} \right)^\delta - 1 \right] \\
 &= \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{j=1}^J p_{ij.} \left[ \left( \frac{p_{ijk}}{p_{i..} p_{.j.}} \right)^\delta - 1 \right] + \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{k=1}^K p_{i.k} \left[ \left( \frac{p_{i.k}}{p_{i..} p_{..k}} \right)^\delta - 1 \right] \\
 &= \frac{2n}{\delta(\delta+1)} \sum_{j=1}^J \sum_{k=1}^K p_{.jk} \left[ \left( \frac{p_{.jk}}{p_{.j.} p_{..k}} \right)^\delta - 1 \right] + \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \left[ \left( \frac{p_{ijk}}{\alpha p_{i..} p_{.j.} p_{..k}} \right)^\delta - 1 \right] \\
 &= CR_{IJ}(\delta) + \boxed{CR_{IK}(\delta)} + CR_{JK}(\delta) + CR_{IJK}(\delta)
 \end{aligned}$$

Cressie-Read family of divergence statistics for only the **row** and **tube** variables (aggregating across the **column** variable)

## Partitioning the Family of Divergence Statistics

Lombardo & Beh (2022) showed that the three-way Cressie-Read divergence statistic can be partitioned as follows

$$\begin{aligned}
 CR(\delta) &= \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \left[ \left( \frac{p_{ijk}}{p_{i..} p_{.j.} p_{..k}} \right)^\delta - 1 \right] \\
 &= \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{j=1}^J p_{ij.} \left[ \left( \frac{p_{ijk}}{p_{i..} p_{.j.}} \right)^\delta - 1 \right] + \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{k=1}^K p_{i.k} \left[ \left( \frac{p_{i.k}}{p_{i..} p_{..k}} \right)^\delta - 1 \right] \\
 &= \frac{2n}{\delta(\delta+1)} \sum_{j=1}^J \sum_{k=1}^K p_{.jk} \left[ \left( \frac{p_{.jk}}{p_{.j.} p_{..k}} \right)^\delta - 1 \right] + \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \left[ \left( \frac{p_{ijk}}{\alpha p_{i..} p_{.j.} p_{..k}} \right)^\delta - 1 \right] \\
 &= CR_{IJ}(\delta) + CR_{IK}(\delta) + \boxed{CR_{JK}(\delta)} + CR_{IJK}(\delta)
 \end{aligned}$$

Cressie-Read family of divergence statistics for only the **column** and **tube** variables (aggregating across the **row** variable)

## Partitioning the Family of Divergence Statistics

Lombardo & Beh (2022) showed that the three-way Cressie-Read divergence statistic can be partitioned as follows

$$\begin{aligned}
 CR(\delta) &= \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \left[ \left( \frac{p_{ijk}}{p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}} \right)^\delta - 1 \right] \\
 &= \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{j=1}^J p_{ij\bullet} \left[ \left( \frac{p_{ijk}}{p_{i\bullet\bullet} p_{\bullet j\bullet}} \right)^\delta - 1 \right] + \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{k=1}^K p_{i\bullet k} \left[ \left( \frac{p_{ijk}}{p_{i\bullet\bullet} p_{\bullet\bullet k}} \right)^\delta - 1 \right] \\
 &= \frac{2n}{\delta(\delta+1)} \sum_{j=1}^J \sum_{k=1}^K p_{\bullet jk} \left[ \left( \frac{p_{ijk}}{p_{\bullet j\bullet} p_{\bullet\bullet k}} \right)^\delta - 1 \right] + \frac{2n}{\delta(\delta+1)} \sum_{i=1}^I \sum_{j=1}^J \sum_{k=1}^K p_{ijk} \left[ \left( \frac{p_{ijk}}{\alpha p_{i\bullet\bullet} p_{\bullet j\bullet} p_{\bullet\bullet k}} \right)^\delta - 1 \right] \\
 &= CR_{IJ}(\delta) + CR_{IK}(\delta) + CR_{JK}(\delta) + CR_{IJK}(\delta)
 \end{aligned}$$

Cressie-Read family of divergence statistics for all **three** the variables (jointly)

## Partitioning the Family of Divergence Statistics

### Example 1:

When  $\delta = 1$

$$CR(1) = CR_{IJ}(1) + CR_{IK}(1) + CR_{JK}(1) + CR_{IJK}(1)$$

is the same as

$$X^2 = X_{IJ}^2 + X_{IK}^2 + X_{JK}^2 + X_{IJK}^2 \quad (\text{see Part III slide 29})$$

### Example 2:

When  $\delta = 0$

$$CR(0) = CR_{IJ}(0) + CR_{IK}(0) + CR_{JK}(0) + CR_{IJK}(0)$$

is the partition of the three-way *likelihood ratio* statistic such that

$$G^2 = G_{IJ}^2 + G_{IK}^2 + G_{JK}^2 + G_{IJK}^2$$

## Partitioning the Family of Divergence Statistics

### Example 1:

When  $\delta = 1$

$$CR(1) = CR_{IJ}(1) + CR_{IK}(1) + CR_{JK}(1) + CR_{IJK}(1)$$

is the same as

$$X^2 = X_{IJ}^2 + X_{IK}^2 + X_{JK}^2 + X_{IJK}^2 \quad (\text{see Part III slide 29})$$

### Example 3:

When  $\delta = 1/2$

$$CR\left(\frac{1}{2}\right) = CR_{IJ}\left(\frac{1}{2}\right) + CR_{IK}\left(\frac{1}{2}\right) + CR_{JK}\left(\frac{1}{2}\right) + CR_{IJK}\left(\frac{1}{2}\right)$$

is the partition of the three-way *Freeman-Tukey* statistic such that

$$T^2 = T_{IJ}^2 + T_{IK}^2 + T_{JK}^2 + T_{IJK}^2$$

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## Partitioning the Family of Divergence Statistics

### Example 1:

When  $\delta = 1$

$$CR(1) = CR_{IJ}(1) + CR_{IK}(1) + CR_{JK}(1) + CR_{IJK}(1)$$

is the same as

$$X^2 = X_{IJ}^2 + X_{IK}^2 + X_{JK}^2 + X_{IJK}^2 \quad (\text{see Part III slide 29})$$

### Example 4:

When  $\delta = 2/3$

$$CR\left(\frac{2}{3}\right) = CR_{IJ}\left(\frac{2}{3}\right) + CR_{IK}\left(\frac{2}{3}\right) + CR_{JK}\left(\frac{2}{3}\right) + CR_{IJK}\left(\frac{2}{3}\right)$$

is the partition of the three-way *Cressie-Read* statistic such that

$$CR = CR_{IJ} + CR_{IK} + CR_{JK} + CR_{IJK}$$

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### Example: Sustainable Development Goals Data

The cross-classification of the three categorical variables (RES, BCA & GEO) that produce the  $4 \times 4 \times 6$  contingency table . . .

Chi-squared test of independence:

- $X^2 = 290.035$
- A p-value  $< 0.001$
- Thus, a statistically significant association exists between at least two of the variables

	RES	BCA			
		(0, 8.13]	(8.13, 23.1]	(23.1, 49.7]	(49.7, 96.4]
<b>Africa</b>		BCA1	BCA2	BCA3	BCA4
(0, 34.5] RES1	1	0	4	8	
(34.5, 53.7] RES2	0	2	4	19	
(53.7, 80.7] RES3	3	2	1	4	
(80.7, 100] RES4	1	1	0	1	
<b>America</b>					
RES1	1	0	6	2	
RES2	0	2	0	1	
RES3	0	3	1	0	
RES4	1	3	2	1	
<b>Asia</b>					
RES1	5	1	2	0	
RES2	6	2	3	1	
RES3	5	3	3	3	
RES4	4	2	0	0	
<b>Australia</b>					
RES1	1	0	0	0	
RES2	0	1	2	0	
RES3	1	1	5	0	
RES4	1	0	0	0	
<b>Carribean</b>					
RES1	3	2	0	0	
RES2	1	1	0	1	
RES3	0	0	0	0	
RES4	2	3	0	0	
<b>Europe</b>					
RES1	1	0	0	0	
RES2	0	0	1	0	
RES3	0	7	5	0	
RES4	3	11	7	3	

### Example: Sustainable Development Goals Data

Partitioning Pearson's Three-way Statistic

Component	$X^2_{IJ}$	$X^2_{IK}$	$X^2_{JK}$	$X^2_{IJK}$	$X^2$
Term	34.970	82.816	95.677	76.573	290.035
P-value	<0.001	<0.001	<0.001	0.002	<0.001
% Contribution	12%	29%	33%	26%	100%

- There is a statistically significant association between at least two of the variables
- There is a statistically significant association between each **pair** of the three variables
- There is a statistically significant association between **all** three variables . . . but . . .
- The association is dominated by the association between BCA & GEO (33%) followed by RES & BCA (29%)

### Example: Sustainable Development Goals Data

The cross-classification of the three categorical variables (RES, BCA & GEO) that produce the  $4 \times 4 \times 6$  contingency table . . .

Chi-squared test of independence:

- $X^2 = 290.035$
- A p-value  $< 0.001$
- Thus, a statistically significant association exists between at least two of the variables

. . . but . . .

- **Sparse contingency table**
- **Multiple zero cell frequencies**

	RES	BCA			
		(0, 8.13] BCA1	(8.13, 23.1] BCA2	(23.1, 49.7] BCA3	(49.7, 96.4] BCA4
<b>Africa</b>					
(0, 34.5] RES1	1	0	4	8	
(34.5, 53.7] RES2	0	2	4	19	
(53.7, 80.7] RES3	3	2	1	4	
(80.7, 100] RES4	1	1	0	1	
<b>America</b>					
RES1	1	0	6	2	
RES2	0	2	0	1	
RES3	0	3	1	0	
RES4	1	3	2	1	
<b>Asia</b>					
RES1	5	1	2	0	
RES2	6	2	3	1	
RES3	5	3	3	3	
RES4	4	2	0	0	
<b>Australia</b>					
RES1	1	0	0	0	
RES2	0	1	2	0	
RES3	1	1	5	0	
RES4	1	0	0	0	
<b>Carribean</b>					
RES1	3	2	0	0	
RES2	1	1	0	1	
RES3	0	0	0	0	
RES4	2	3	0	0	
<b>Europe</b>					
RES1	1	0	0	0	
RES2	0	0	1	0	
RES3	0	7	5	0	
RES4	3	11	7	3	

### Example: Sustainable Development Goals Data

Partitioning the Three-way Freeman-Tukey Statistic

Component	$T_{ij}^2$	$T_{jk}^2$	$T_{ik}^2$	$T_{ijk}^2$	$T^2$
Term	35.112	83.340	92.421	39.619	250.491
P-value	<0.001	<0.001	<0.001	0.699	<0.001
% Contribution	14%	33%	37%	16%	100%

- There is a statistically significant association between at least two of the variables
- There is a statistically significant association between each **pair** of the three variables
- . . . but . . . the three-way association is **not** statistically significant (p-value = 0.699)
- The association is still dominated by the association between BCA & GEO (37%) followed by RES & BCA (33%)

**Example: Sustainable Development Goals Data**

Multi-Way Correspondence Analysis

Partitioning the Three-way Cressie-Read Statistic

Component	CR <sub>Ij</sub>	CR <sub>IK</sub>	CR <sub>Jk</sub>	CR <sub>IJK</sub>	CR
Term	34.998	82.653	93.105	48.559	259.035
P-value	<0.001	<0.001	<0.001	0.332	<0.001
% Contribution	13%	32%	36%	19%	100%

- There is a statistically significant association between at least two of the variables
- There is a statistically significant association between each **pair** of the three variables
- . . . but . . .the three-way association is **not** statistically significant (p-value = 0.332)

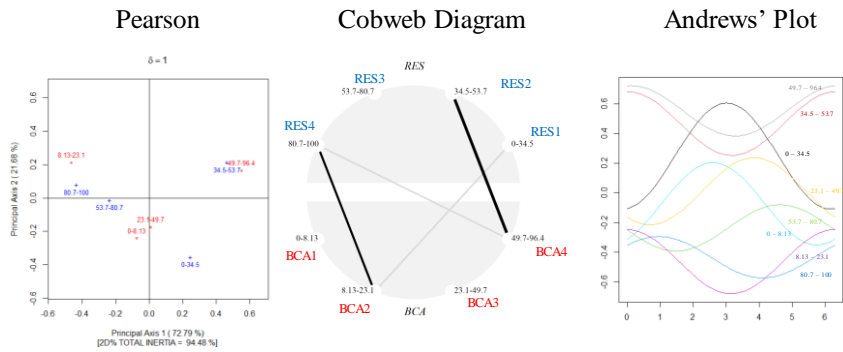
• The association is dominated by the association between BCA & GEO (36%) followed by RES & BCA (32%)

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**Example: Sustainable Development Goals Data**

Multi-Way Correspondence Analysis

RES vs BCA



$$X_{Ij}^2 = 34.97$$

(12% of total inertia)

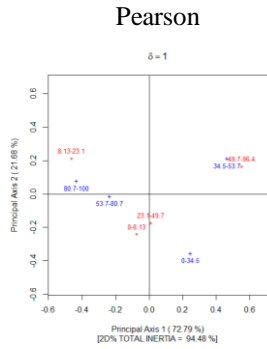
60



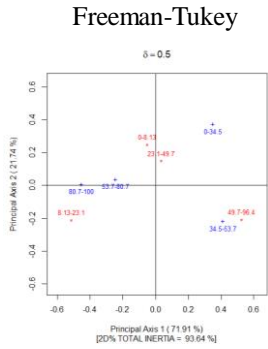
**Example: Sustainable Development Goals Data**

Multi-Way Correspondence Analysis

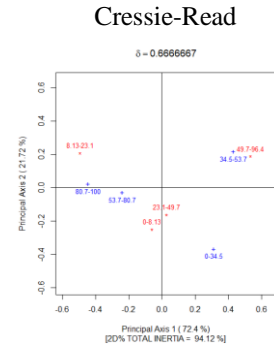
RES vs BCA



$X^2_{IJ} = 34.97$   
(12% of total inertia)



$T^2_{IJ} = 35.112$   
(14% of total inertia)



$CR_{IJ} = 34.998$   
(13% of total inertia)

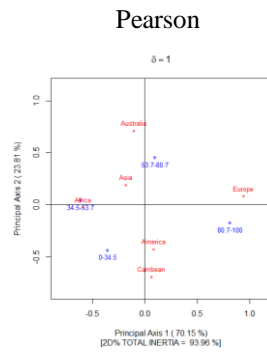


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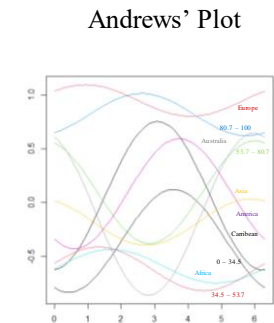
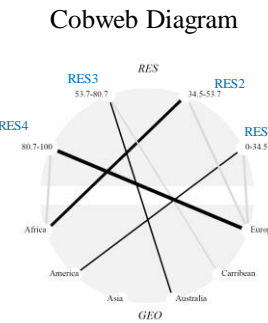
**Example: Sustainable Development Goals Data**

Multi-Way Correspondence Analysis

RES vs GEO



$X^2_{IK} = 82.816$   
(29% of total inertia)

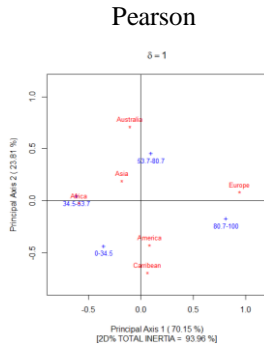


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**Example: Sustainable Development Goals Data**

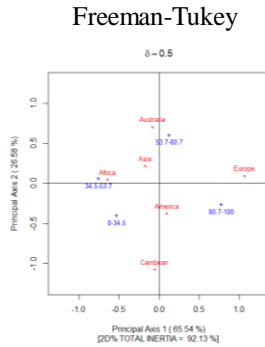
Multi-Way Correspondence Analysis

RES vs GEO



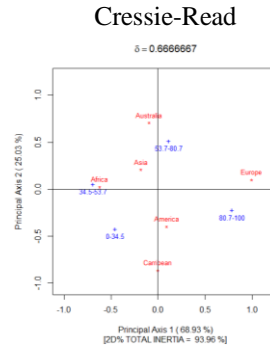
$$X^2_{JK} = 82.816$$

(29% of total inertia)



$$T^2_{JK} = 83.340$$

(35% of total inertia)



$$CR_{JK} = 82.653$$

(32% of total inertia)

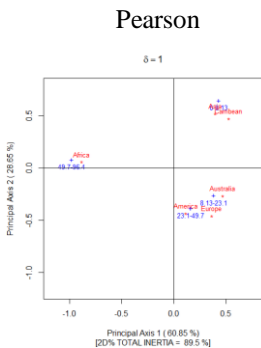


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**Example: Sustainable Development Goals Data**

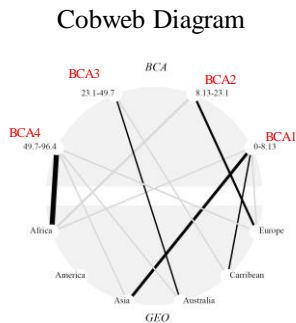
Multi-Way Correspondence Analysis

BCA vs GEO

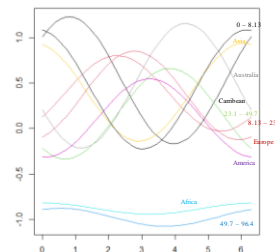


$$X^2_{JK} = 95.677$$

(33% of total inertia)



Andrews' Plot

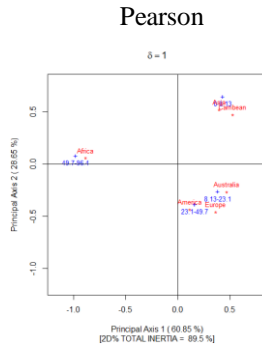


64



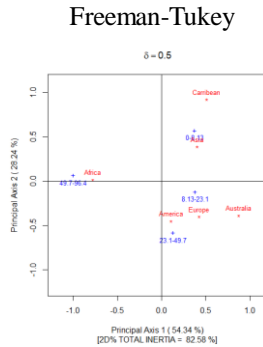
## Example: Sustainable Development Goals Data

### BCA vs GEO



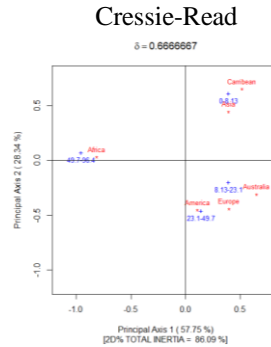
$$X_{JK}^2 = 95.677$$

(33% of total inertia)



$$T_{JK}^2 = 92.421$$

(37% of total inertia)



$$CR_{JK} = 93.105$$

(36% of total inertia)



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## Appendix A: Indicator Matrix

```

indicator.exe <- function(N) {

#####
# Converts an m-way contingency table, N, into its indicator matrix form #
#####

nvars <- dim(N)                                     # number of variables cross-classified to form N
ncats <- length(nvars)                             # number of categories
catnames <- NULL
for(l in 1:ncats) {
  catnames <- c(catnames, dimnames(N)[[l]])
}
b <- c()
for(ii in 1:ncats) {
  aaap <- prod(nvars[(ii + 1):ncats])
  paaa <- prod(nvars[1:(ii - 1)])
  if(ii == 1) paaa <- 1
  if(ii == ncats) aaap <- 1
  b <- cbind(b, rep(rep(1:nvars[ii]), rep(aaap, nvars[ii])), paaa))
}
bb <- matrix(0, dim(b)[1], sum(nvars))
for(jj in 1:dim(b)[1]) {
  eee <- c(b[jj, 1])
  for(ii in 2:length(nvars))
    eee <- c(eee, eval(parse(text = paste("b[jj, ", ii, "] + sum(nvars[seq(1, ii-1)])", sep = ""))))
  bb[jj, eee] <- 1
}
Z <- c()
for(jj in 1:dim(b)[1]) {
  aaaa <- c()
  for(ii in 1:dim(b)[2]) {
    aaaa <- paste(aaaa, b[jj, ii], sep = ",")
  }
  bbbb <- eval(parse(text = paste("N[, substring(aaaa, 2), "]", sep = "")))
  if (bbbb > 0) {
    Z <- rbind(Z, matrix(rep(bb[jj, ]), bbbb), bbbb, byrow = T)
  }
}
dimnames(Z) <- list(paste("Individual", 1:sum(N)), paste(catnames))
Z
# The indicator matrix of size n x ncats
}

```



## Appendix B: Burt Matrix

R Code

```
burt.exe <- function(N) {  
#####  
#       Converts an m-way contingency table, N, into its Burt matrix form       #  
#####  
  
    Z <- indicator.exe(N)  
    Burt <- as.matrix(t(Z) %*% Z)  
    Burt  
}
```

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